### EC408 Topics in Applied Econometrics

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The F test is a very useful test statistic for assessing the importance of a group of variables.

- Given two models, one with and one without the group of variables to be tested
  - it calculates the change in the level of explanation per parameter (or degree of freedom)
  - and compares it with the amount of random variation per degree of freedom in the model.
  - For the group of variables to have had a significant effect, we should expect the change in explanation per additional degree of freedom to exceed the amount of random variation per degree of freedom in the model.

$$\frac{D_1 - D_2}{k'}$$
$$\frac{D_2}{T - k}$$

 $F = \frac{\frac{D_1 - D_2}{k'}}{\frac{D_2}{T - k}}$ 

- $D_1$  is the unexplained variation (the residual sum of squares) for model 1
- $D_2$  is the residual sum of squares for model 2
- model 2 is the same as model 1 plus the additional k' explanatory variables (the lags to be tested)
- then D<sub>1</sub> > D<sub>2</sub> and the additional explanation is the reduction in RSS, and per degree of freedom lost this is

$$\frac{D_1 - D_2}{k'} \tag{61}$$

• This quantity is compared with the random variation per model degree of freedom (T - k), given by

$$\frac{D_2}{T-k} \tag{62}$$

- Assuming that  $D_2$  is calculated from white noise residuals.
- Assuming that the additional k' explanatory variables (lags) are in fact null (have no effect), then the ratio

$$F = \frac{\frac{D_1 - D_2}{k'}}{\frac{D_2}{T - k}} \sim F_{k', T - k}$$
(63)

### **F-distribution**



The F-distribution is simply the ratio of two chi-squared
variables divided by their respective degrees of freedom i.e.

$$\frac{C_1/d_1}{C_2/d_2} = F(d_1, d_2) \qquad 0 \le F \le \infty$$



Abridged from M. Merrington and C. M. Thompson, 'Tables of percentage points of the inverted beta (F) distribution', *Biometrika*, vol. 33, 1943, p. 73. By permission of the *Biometrika* trustees.

 $v_2$  = degrees of freedom for denominator

f(F)

We should start with a complex model with several lags, which will tend to produce white noise residuals which are then appropriate for hypothesis testing (remember if we have autocorrelation or heteroscedasticity in the errors then standard errors and inference will be biased). This is in line with the top-down philosophy of Hendry.

The first test lag m conditional on the presence of all other lags

It asks the question, does lag m carry any explanatory power given the existence of the other lags in the model?

The second test is of the significance of omitted lags up to lag m

Hence we test for 1...m, then 2...m, then 3...m, etc

The first test lag m conditional on the presence of all other lags

. Given the model

$$Y_{t} = \alpha_{1}Y_{t-1} + b_{10}X_{1t} + b_{11}X_{1t-1} + b_{12}X_{1t-2} + b_{13}X_{1t-3} + b_{14}X_{1t-4} + b_{20}X_{2t} + b_{21}X_{2t-1} + b_{22}X_{2t-2} + b_{23}X_{2t-3} + b_{24}X_{1t-4} + u_{t}$$
(64)

There are k' = 3 variables for m = 1,

 $Y_{t-1}$  or CONS\_1,  $X_{1t-1}$  or INC\_1, and  $X_{2t-1}$  or INFLAT\_1,

The denominator of the F ratio is  $D_2/(T-k) = \frac{177.365821}{(155-11)}$  for the 4

lag unrestricted model (4 lags to create white noise residuals) with degrees of freedom T - k = 155-11 = 144, since there are 155 observations (omitting 4 to allow 4 lags) and 11 parameters, as shown by counting the variables in the model.

Evidently lag 1 has significant explanatory power, since F = 172.54 has a p-value close to zero in the theoretical  $F_{3,144}$  distribution. The test statistic is very atypical of the reference distribution assumed under the null.

### EQ(15) Modelling CONS by OLS (using data.in7)

The estimation sample is: 1954 (1) to 1992 (3)

	Coefficient	Std.Error	t-value	t-prob	Part.R <sup>2</sup>
CONS_1	0.811974	0.03601	22.6	0.000	0.7793
INC	0.478975	0.02917	16.4	0.000	0.6518
INC_1	-0.291088	0.04081	-7.13	0.000	0.2611
INC_2	-0.0362995	0.03775	-0.961	0.338	0.0064
INC_3	0.0400228	0.03782	1.06	0.292	0.0077
INC_4	-0.00484922	0.03227	-0.150	0.881	0.0002
INFLAT	-0.910971	0.2631	-3.46	0.001	0.0768
INFLAT_1	0.0640396	0.4672	0.137	0.891	0.0001
INFLAT_2	-0.0435876	0.4845	-0.0900	0.928	0.0001
INFLAT_3	-0.316838	0.4554	-0.696	0.488	0.0034
INFLAT_4	0.112654	0.2722	0.414	0.680	0.0012
sigma	1.10982	RSS		<mark>177.3658</mark>	<mark>321</mark>
log-likelihood	<mark>-230.382</mark>	DW		1.	.91
no. of observation	ເຣ 155	no. of par	ameters		11
mean(CONS)	875.653	var(CONS) 1			<b>182</b>

### Tests on the significance of each lag

Lag	1	F(3,144)	=	172.54	[0.0000]**
Lag	2	F(2,144)	=	0.46687	[0.6279]
Lag	3	F(2,144)	=	0.78841	[0.4565]
Lag	4	F(2,144)	=	0.086535	[0.9172]

Clearly the additional explanatory power of lag 2 (k' = 2 additional variables), is weak, given that we have the other lags 1,3, and 4 present in the model. The contribution of lag 2 does not differ much from random variation per degree of freedom (the denominator of the F ratio). Likewise, lags 3 and 4 carry no additional explanatory power.

The second test is of the significance of omitted lags up to lag m

$$Y_{t} = \alpha_{1}Y_{t-1} + b_{10}X_{1t} + b_{11}X_{1t-1} + b_{12}X_{1t-2} + b_{13}X_{1t-3} + b_{14}X_{1t-4} + b_{20}X_{2t} + b_{21}X_{2t-1} + b_{22}X_{2t-2} + b_{23}X_{2t-3} + b_{24}X_{1t-3} + u_{t}$$
(64)

We commence with the set of lags 1... 4 inclusive.

- Looking at the four lag model that this involves 9 variables, so the appropriate degrees of freedom for the F test are k' = 9 and T k = 144.
- The test statistic 173.84 has a very low p-value in the reference distribution under the null, which is the  $F_{k',T-k}$ . At least one lag in lags 1...4 is significant.

Lag	1	-	4	F(9,144)	=	173.84	[0.0000]**
Lag	2	-	4	F(6,144)	=	<mark>0.46643</mark>	[0.8323]
Lag	3	-	4	F(4,144)	=	0.44169	[0.7783]
Lag	4	-	4	F(2, 144)	=	0.086535	[0.9172]

Next we test lags 2...4, which involves k' = 6 variables and T - k = 144. This has a high p-value so the test statistic 0.46643 is typical of the reference distribution  $F_{k',T-k}$  indicating that we should not reject the null that the effects of lags 2...4 are no more than random variation. Similarly, testing lags 3...4 and lag 4 retains the null. It appears from this analysis that the first lag alone is significant.

an entirely equivalent procedure can be accomplished by looking at the log likelihoods of nested models

where we have the log likelihoods of two models we wish to compare, a restricted model giving log  $L_R$  and an unrestricted model giving log  $L_U$ 

Under the null that the k restrictions are true, then a good approximation is

$$-2\log\left(\frac{L_R}{L_U}\right) = 2\log\left(\frac{L_U}{L_R}\right) = 2\left\{\log L_U - \log L_R\right\} \sim \chi_k^2$$
(65)

We check the whether the likelihood ratio test also does not reject the null that lags 2...4 are zero.

If we compare the 4 lag model (log  $L_U$ ) with the 1 lag model (log  $L_R$ ),

log  $L_U = -230.382$  (it is unrestricted because no restriction have been placed on the parameters, so all 4 lags are present),

log  $L_R = -231.873$  for the 1 lag model (entailing 6 parameters restricted to zero).

Refer 2{ -230.382 +231.873 } = 2.982 to the  $\chi_6^2$  distribution (there are 2 variables at lag 2, 2 at lag 3 and 2 at lag 4).

It has a p-value equal to 0.8111, again entirely equivalent to the outcome of the F test. We do not reject the null that lags 2...4 are zero.

comparing the 4 lag model with the zero lag model tests whether at least one of lags 1...4 is significant.

 $\log L_U = -230.382$ 

 $\log L_{R} = -422.084$ 

the test statistic is  $2\{-230.382 + 422.084\} = 383.404$ has a p-value close to 0 in  $\chi_9^2$